Applications of Astrostatistics to Dark Matter Phenomenology and Beyond the Standard Model Theories

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Evidence for dark matter

 Multiple astrophysical and cosmological probes strongly point to the existence of non-baryonic dark matter:



Gravitational lensing M/L >> 1

Kent 1987

Evidence for dark matter (cont'ed)

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The cosmic microwave background

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Change in the matter/radiation energy density ratio at decoupling:

Combined cosmological probes

SUSY in a nutshell

- Every SM particle acquires a superpartner with the same quantum numbers but opposite spin statistics
- Unbroken SUSY: sparticles have the same mass as SM particles
- SUSY must be broken

Standard Model particles and fields

I second a second l				re percenter		
		Interactio	on eigenstates	Mass eig	enstates	
Symbol	Name	Symbol	Name	Symbol	Name	
q=d,c,b,u,s,t	quark	\tilde{q}_L,\tilde{q}_R	squark	\tilde{q}_1,\tilde{q}_2	squark	
$l = e, \mu, \tau$	lepton	$ ilde{l}_L, ilde{l}_R$	$_{\rm slepton}$	\tilde{l}_1,\tilde{l}_2	slepton	
$\nu = \nu_e, \nu_\mu, \nu_\tau$	neutrino	$\tilde{\nu}$	$\operatorname{sneutrino}$	$\tilde{ u}$	$\operatorname{sneutrino}$	
g	gluon	\tilde{g}_{\perp}	gluino	\widetilde{g}	gluino	
W^{\pm}	W-boson	\tilde{W}^{\pm}	wino			
H^-	Higgs boson	\tilde{H}_1^-	higgsino	$\tilde{\chi}_{1,2}^{\pm}$	chargino	
H^+	Higgs boson	\tilde{H}_2^+	higgsino	,		
В	B-field	\tilde{B}	bino j			A natural
W^3	W^3 -field	\tilde{W}^3	wino			
H_{1}^{0}	Higgs boson	$\tilde{\mathbf{u}}_{0}$	higgsing }	$\tilde{\chi}^{0}_{1,2,3,4}$	neutralino	DM
H_{2}^{0}	Higgs boson	\tilde{n}_1	niggsino			
$H_3^{ar 0}$	Higgs boson	H_{2}^{0}	higgsino)			candidate

Supersymmetric partners

Review article: Bertone, Hooper & Silk (2005)

A solution to the DM problem

- Under R-parity, the neutralino is stable
- Neutral, weakly interacting, mass in the range ~ 100 GeV to a few TeV
- Cosmological neutralino relic abundance:

$$\Omega_{\chi}h^2 \sim \frac{3\times 10^{-27} {\rm cm}^3 {\rm s}^{-1}}{\langle \sigma v \rangle}$$

Courtesy Sabine Kraml

The model & data

- The general Minimal Supersymmetric Standard Model (MSSM): 105 free parameters!
- Need some (pretty strong) simplifying assumption: the Constrained MSSM (CMSSM) reduces the free parameters to just 4 continous variables plus a discrete one (sign(µ)).
- Clearly a highly constrained model (probably not the end of the story!)
- Present-day data: collider measurements of rare processes, CDM abundance (Planck), sparticle masses lower limits, EW precision measurements, Higgs mass and couplings.
- Astrophysical direct and indirect detection techniques might also be competitive: neutrino (IceCUBE), gamma-rays (Fermi), antimatter (PAMELA), direct detection (XENON1T, LUX, PandaX,...)

- Points accepted/rejected in a in/out fashion (e.g., 2-sigma cuts)
- No statistical measure attached to density of points: no probabilistic interpretation of results possible
- Inefficient/Unfeasible in high dimensional parameters spaces (N>3)
- Explores only a very limited portion of the parameter space!

2 dimensional slices

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Roszkowski et al (2001)

"Global fits": multi-parameters, simultaneous likelihood-based fits to data in several observational channels (direct/indirect detection+colliders+cosmology), often including uncertainties from poorly-known nuisance parameters (e.g., astrophysical quantities)

Why global fits?

- The theoretical interpretation of a Dark Matter-like signal requires fitting an underlying model (e.g., SUSY, extra dimensions, etc) to the data.
- In case of a detection, a global multi-messenger approach will allow to check the consistency of the theory across observables and to obtain more stringent constraints on the Dark Matter properties.
- Robust and believable interpretation of direct and indirect detection data requires a careful modeling of astrophysical and experimental uncertainties.

Global CMSSM scans

- Bayesian approach introduced by two groups (early work by Baltz & Gondolo, 2004):
- Ben Allanach (DAMPT) and collaborators (Allanach & Lester, 2006 onwards)
- Ruiz de Austri, Roszkowski & RT (Ruiz de Austri et al, 2006 onwards)
 + Feroz & Hobson (MultiNest), + Silk (indirect detection), + Strigari (direct detection), + Martinez et al (dwarfs), + de los Heros (IceCube)

See also: Ellis et al (2004 onwards), Buchmuller et al (2008, 2009), Scott et al (2009), Akrami et al (2009)

Generic analysis pipeline for BSM physics Imperial College

Analysis pipeline for BSM physics

Analysis pipeline for BSM physics

 10^{3}

Exploration with "random scans"

- Points accepted/rejected in a in/out fashion (e.g., 2-sigma cuts)
- No statistical measure attached to density of points: no probabilistic interpretation of results possible, although the temptation cannot be resisted...
- Inefficient in moderately large dimensional parameters spaces (even just D>5)

- E.g.: Fermi constraints on DM gamma-ray flux from dwarfs compared with theory
- MSSM UMa II ····· Draco WMAP compatible Coma Berenices ····· Sextans 10⁵ below WMAP ····· Fornax (σ **v**) 95%^{C.L.} 100% b b Bootes 10⁴ (10⁻²⁶ cm³s⁻¹) 10⁻²⁶ cm³s⁻¹) **\$**0**\$** 10⁻¹

m_{DM} (GeV)

10²

Abdo et al, arxiv: 1001.4531

10⁻²

10⁻³

Statistical inference

- Given a model *M*, with parameters of interest θ, nuisance parameters Ψ, a prior pdf P(θ, Ψ | M), and available data *d* with likelihood P(d | θ, Ψ, M) = L(θ, Ψ), we need an algorithm to compute the following statistical quantities:
- **1. Marginal posterior pdf:**

$$P(\theta|d, M) = \int d\Psi P(\theta, \Psi|d, M) = \int d\Psi \frac{L(\theta, \Psi) P(\theta, \Psi|M)}{P(d|M)}$$

2. Profile likelihood ratio:

$$\lambda(\theta) = \frac{L(\theta, \hat{\psi})}{L(\hat{\theta}, \hat{\psi})}$$

where $\hat{\psi}$ is the conditional MLE and $\hat{\psi}$ the unconditional MLE

3. Bayesian evidence (model likelihood):

$$P(d|M) = \int d\theta d\Psi L(\theta, \Psi) P(\theta, \Psi|M)$$

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Solution: global fits

Carry out a **simultaneous fit** of all relevant SUSY and SM parameter to the experimental data/constraints.

Marginalize (= integrate) or maximise along the hidden dimensions to obtain results that account for the multidimensional nature of the problem.

Global Fits: Some History

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2D 95% CL Startvalues

1D 68% CL Startvalues 1

2D 95% CL Startvalues 2

1D 68% CL Startvalues 2

600

GAMBIT 1.0.0

700

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8000 10000

First Generation Global Fits: SuperBayeS

SuperBayeS Supersymmetry Parameters Extraction Routines for Bayesian Statistics

- Implements the CMSSM, but can be easily extended to the general MSSM
- Latest release (v 1.5.1) in April 2011: linked to SoftSusy 2.0.18, DarkSusy 5.0, MICROMEGAS 2.2, FeynHiggs 2.5.1, Hdecay 3.102. Uses MultiNest v 2.8.
- Includes up-to-date constraints from all observables, plotting routines, statistical analysis tools, posterior and profile likelihood plots. Fully parallelized, MPI-ready, user-friendly interface
- MCMC engine (Metropolis-Hastings, bank sampler), grid scan mode, multi-modal nested sampling MultiNest algorithm (Feroz & Hobson 2008)
 A full 8D scan now takes less than 2 days on 8 CPUs.
- Now discontinued

Second Generation Global Fits: GAMBIT

GAMBIT: The Global And Modular BSM Inference Tool

gambit.hepforge.org

EPJC **77** (2017) 784

arXiv:1705.07908

- Extensive model database not just SUSY
- Extensive observable/data libraries
- Many statistical and scanning options (Bayesian & frequentist)
- Fast LHC likelihood calculator
- Massively parallel
- Fully open-source

 Members of: ATLAS, Belle-II, CMS, CTA, Fermi-LAT, DARWIN, IceCube, LHCb, SHiP, XENON
 Authors of: DarkSUSY, DDCalc, Diver, FlexibleSUSY, gamlike, GM2Cale

DarkSUSY, DDCalc, Diver, FlexibleSUSY, gamlike, GM2Calc, IsaJet, nulike, PolyChord, Rivet, SOFTSUSY, SuperIso, SUSY-AI, WIMPSim

- Fast definition of new datasets and theories
- Plug and play scanning, physics and likelihood packages

Collaborators:

Peter Athron, Csaba Balázs, Ankit Beniwal, Florian Bernlochner, Sanjay Bloor, Torsten Bringmann, Andy Buckley, Eliel Camargo-Molina, Marcin Chrząszc, Jan Conrad, Jonathan Cornell, Matthias Danninger, Tom Edwards, Joakim Edsjö, Ben Farmer, Andrew Fowlie, Tomás Gonzalo, Will Handley, Sebastian Hoof, Selim Hotinli, Felix Kahlhoefer, Suraj Krishnamurthy, Anders Kvellestad, Julia Harz, Paul Jackson, Tong Li, Greg Martinez, Nazilla Mahmoudi, James McKay, Are Raklev, Janina Renk, Chris Rogan, Roberto Ruiz de Austri, Patrick Stoecker, Roberto Trotta, Pat Scott, Nicola Serra, Daniel Steiner, Puwen Sun, Aaron Vincent, Christoph Weniger, Sebastian Wild, Martin White, Yang Zhang

40+ participants in 10 Experiments & 14 major theory codes

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The Global And Modular BSM Inference Tool

- A new framework for BSM global fits
- Fully open source
- · Modular design: easily extended with
 - new models
 - new likelihoods
 - new theory calculators
 - new scanning algorithms
- · Use external codes (backends) as runtime plugins
 - Currently supported:
 - C, C++, Fortran, Mathematica
 - Coming soon: Python
- · Two-level parallellization with MPI and OpenMP
- · Hierarchical model database
- · Flexible output streams (ASCII, HDF5, ...)
- · Many scanners and backends already included

CMSSM: Frequentist Fits

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SuperBayeS: profile likelihood

2011, 41M samples (~1 week on 8 CPUs) incl. ATLAS 35 fb⁻¹ Run I and WMAP-7

GAMBIT: profile likelihood (95% CL)

2018, 71M samples (3 days on 2400 CPUs) incl. ATLAS/CMS Run I + II, Planck 2018

Stau co-annihilation region now ruled out at > 95% CL in the CMSSM.

CMSSM: Bayesian posteriors (2011)

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Tentative convergence between Frequentist and Bayesian scans has not improved despite 7 years of additional data (and stronger LHC lower limits). The relative viability of the Focus Point region cannot be robustly established

Direct Detection Prospects

Profile likelihood, 2011

Astrophysical parameters fixed Relic density constraint imposed

Profile likelihood, 2018

Astrophysical parameters varied Relic density as upper limit only

The importance of modeling the Milky Way

- Assuming an incorrect local density (by a factor of 2) can lead to a 15 sigma bias in the reconstructed cross section
- Milky Way modeled as a parameterized bulge+halo+disk spherical, isotropic superposition.
 Sloan-like data used to constrain the model's parameters
- Marginalization over Milky Way halo model parameters converts catastrophic direct detection systematic errors into more manageable statistical errors

Parameterized Milky Way model (7 parameters)

Strigari & Trotta (0906.5361)

Log scattering cross section (pb)

Identification of Cosmological DM

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If a signal is seen both at the LHC and in direct detection detectors, how can we check that this WIMP makes up the bulk of the cosmological relic density?

- Fit low-energy SUSY parameters and try to predict Ωh^2 from LHC data alone.
- **Problem**: LHC data alone are unable to constrain the relic abundance. Even DD data cannot break the degeneracy (if ρ_x assumed fixed):

Complementarity of direct detection and LHC data

• **Strategy**: assume that the local density scales with the cosmological relic abundance ("scaling Ansazt"):

LHC data only (300 fb⁻¹) LHC + DD (fixed ρ_x) Bertone et al. (2010) Bertone et al. (2010) Bertone et al. (2010) LHC + 1 ton DD (fixed p) LHC + 1 ton DD (scaling p_) LHC only $\text{log}_{10}(\sigma_{\chi-p}^{\text{SI}}/\text{pb})$ $\log_{10}(\sigma_{\chi-p}^{SI}/pb)$ $\text{log}_{10}(\sigma_{\chi-p}^{\text{SI}}/\text{pb})$ X -10-10 -11-11 -3 -3 -3 -2 -2 -2 -1 0 0 0 -1 $\log_{10}(\Omega_{\chi^0_{\star}} h^2)$ $\log_{10}(\Omega_{\gamma^0} h^2)$ $\log_{10}(\Omega_{\chi^0_{\star}}\,\text{h}^2)$ **Our scaling Ansatz breaks** degeneracy in parameter space Bertone, RT et al, 1005.4280

Cosmological solution identified!

LHC + DD (ρ_x scales with Ωh^2)

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Wedding-cake approach

- A full MSSM-15 profile likelihood scan might require O(10⁹) ATLAS likelihood evaluations
- Not feasible with full (expensive simulations). Wedding-cage strategy:

CPU-time per

- Coverage of intervals is a frequentist property.
- Guaranteed when using e.g. Feldman-Cousins procedure to build intervals.
- Approximate confidence intervals are obtained via the Neyman construction with profile likelihood ratio as a test statistic.
- From a Bayesian perspective, coverage properties of credible intervals (if desired) can be used to calibrate priors.
- Coverage studies are computationally expensive:
 - (a) choose fiducial point in parameter space
 - (b) generate pseudo-data
 - (c) reconstruct credible/confidence interval
 - (d) check whether fiducial point within/without interval.

"Instantaneous" inference with neural networks

 Standard MCMC (SuperBayeS v1.23, 2006 release)
 720 CPU days

MultiNest

 (SuperBayeS v1.5, 2010 release)
 16 CPU days
 speed-up factor: ~ 50

 SuperBayeS+Neural Networks (Bridges, Cranmer, Feroz, Hobson, Ruiz & RT, <u>1011.4306</u>)
 less than 1 CPU minute speed-up factor: 30'000

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Neural nets shortcuts

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Neural networks technology

Bridges, KC, RT et al (<u>1011.4306</u>)

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- We used a feed-forward multi-layer perceptron to "replace" SoftSusy in predicting the weak-scale masses from the CMSSM input parameters
- After training with a few 1000's samples, the neural net achieved a correlation > 99.99%

Coverage: are intervals what they say?

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- We did 10,000 reconstructions, each with 1 million samples. This would have taken 1,100 CPU yrs using standard methods. Neural network speed-up is dramatic, of order 10⁴.
- Test case: use weak-scale masses as input, with Gaussian likelihood. Coverage is exact (within noise), as expected:

Profile likelihood

Coverage: are intervals what they say?

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 Mapping back constraints to the CMSSM parameters, we find substantial overcoverage for both Bayesian and profile likelihood intervals:

Origin of over-coverage in the CMSSM

• The CMSSM prior introduces "physicality" boundaries in the weak-scale masses space. As a consequence, the distribution of -2 ln(λ) is not well approximated by χ^2 and Wilks' theorem does not apply.

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- Global scans of Beyond the Standard Model theories are now a mature field
- Ensuing predictions for direct/indirect detection experiments have to be interpreted with care due to weakness of current constraints
- BUT, quantitative, accurate inference tools will be required to pursue a robust and believable multi-messenger identification of DM
- Current tools have the ability to include statistical and systematic uncertainties in a statistically principled way
- Further work will focus on the careful assessment of their performance on simulated data sets under a variety of conditions

Gamma-ray from WIMP annihilation

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Predicting the gamma ray flux

Differential flux:

particle astrophysics physics

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DM density profile:

$$\rho_{\chi}(r) = \rho_0 \frac{(r/r_0)^{-\gamma}}{[1 + (r/a)^{\alpha}]^{\frac{\beta - \gamma}{\alpha}}} [1 + (r_0/a)^{\alpha}]^{\frac{\beta - \gamma}{\alpha}}$$

Halo model	a~(m kpc)	α	β	γ
isothermal cored	3.5	2	2	0
NFW	20.0	1	3	1
NFW+ac	20.0	0.8	2.7	1.45
Moore	28.0	1.5	3	1.5
Moore+ac	28.0	0.8	2.7	1.65

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Predicting the GC gamma ray signal

Einasto profile:

DM density profile

Halo model	a~(m kpc)	α	eta	γ	$ar{J}(10^{-3}{ m sr})$	$ar{J}(10^{-5}{ m sr})$
isothermal cored	3.5	2	2	0	30.35	30.40
NFW	20.0	1	3	1	1.21×10^3	$1.26 imes 10^4$
NFW+ac	20.0	0.8	2.7	1.45	1.25×10^5	1.02×10^7
Moore	28.0	1.5	3	1.5	1.05×10^5	$9.68 imes 10^6$
Moore+ac	28.0	0.8	2.7	1.65	$1.59 imes 10^6$	$3.12 imes 10^8$

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A GC Excess?

Goodenough & Hooper (2009)

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- Initial hints of a DM signal from the galactic center (Goodenough & Hooper 2009)
- Caveats:
 - GC very complex region: point sources, diffuse emission, etc
 - Strong bounds from PAMELA antiproton data and radio data constraining synchroton emission
 - The Hooper & Goodenough model is effectively ruled out.

GC Excess from the pMSSM

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- Astrophysical explanations most plausible for the "excess" (unresolved population of millisecond pulsars, inverse Compton from cosmic rays).
- Can the putative excess be explained with a SUSY model? YES

Bertone, RT et al 2016

GC Excess from the pMSSM

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- There exist points in the 19-dimensional pMSSM that can explain the GC excess while satisfying all other constraints
- Since this study was done, direct detection constraints have further improved, largely ruling out the ttbar island.

Milky Way dwarf galaxies

Nearby, lots of dark matter ($\log_{10} J \sim 18 - 20$)

Not much else: no astrophysical background* compare to Galactic center

Dwarf searches reach the relic cross section

Hoof, Geringer-Sameth & RT (in prep)

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6.5 yr Pass 7 Fermi LAT data

6.5 yr Pass 8 Fermi LAT data

Bonnivard+ 1504.03309 (ApJL)

Profile likelihood

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6.5 yr Pass 8 Fermi LAT data

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To discover dark matter annihilation using dwarfs:

1. Gamma-ray data is inconsistent with background

2. Inconsistent with any other possible source (e.g. non-DM astrophysics, incorrect diffuse bg models)

3. Consistent with dark matter annihilation (compare with other dwarfs, other experiments)

Two ways to model background give two different significances. Beware of the Look Elsewhere Effect, too!

Diffuse background model

- Poisson with given spectrum:
 - "physical" model cosmic ray interactions in Milky Way, extragalactic isotropic emission, charged particle misidentification
 - No additional non-DM sources along line of sight towards dwarf

 H_0 : No additional source p-value = 0.0001

0.0001 *p-value* = 0.01

Slide: Alex Geringer-Sameth

Geringer-Sameth & Koushiappas 1108.2914 (PRL), Geringer-Sameth+ 1410.2242 (PRD), Geringer-Sameth+ 1503.02320 (PRL)

Empirical background from sampling

To discover dark matter annihilation using dwarfs:

1. Gamma-ray data is inconsistent with background

2. Inconsistent with any other possible source (e.g. non-DM astrophysics, incorrect diffuse bg models)

3. Consistent with dark matter annihilation (compare with other dwarfs, other experiments)

Compare with known classes of gamma-ray sources

Geringer-Sameth+ 1807.08740

Slide: Alex Geringer-Sameth

Combine all three tests into a single search using a Bayesian approach

Carina 2 and 3 Discovered by Torrealba+ 1801.07279 (MNRAS)

Idea: empirical prior distribution of background sources leads to empirical prior on non-DM point-like sources around the target

Conclusions - Constraints on DM from dwarfs

- Ruling out diffuse bg model is not enough
- Want to distinguish DM annihilation from non-DM source populations without sacrificing sensitivity
- Methods apply to any dwarf which is a promising DM target and shows evidence for gamma-ray emission along line of sight
- Should be simple to extend to any dark matter target where you expect localized emission (e.g. galaxy clusters, groups, dark subhalos)

SUPPLEMENTARY MATERIAL

Profile likelihood results: comparison

- Akrami et al (0910.3950) adopted a genetic algorithm (GA) to map out the profile likelihood.
- This allows to find isolated spikes in the likelihood in the focus point region: is this something other frequentist fits might have missed?

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Samples from priors only

· No data in the likelihood, non-physical points discarded

Trotta et al (2008)

Samples from priors only

Flat priors

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Priors are highly informative regarding the quantities being constrained!

Current constraints

Indirect observables

Observable	Mean value Uncertainties		ainties	ref.
	μ	σ (exper.)	τ (theor.)	
M_W	80.398 GeV	$25 { m MeV}$	$15 { m MeV}$	[30]
$\sin^2 heta_{ m eff}$	0.23153	$16 imes 10^{-5}$	$15 imes 10^{-5}$	[30]
$\delta a_{\mu}^{ m SUSY} imes 10^{10}$	29.5	8.8	1.0	[31]
$BR(\overline{B} \to X_s \gamma) \times 10^4$	3.55	0.26	0.21	[32]
ΔM_{B_s}	17.77 ps^{-1}	$0.12 \ {\rm ps}^{-1}$	$2.4 \ \mathrm{ps}^{-1}$	[33]
$BR(\overline{B}_u \to \tau \nu) \times 10^4$	1.32	0.49	0.38	[32]
$\Omega_{\chi}h^2$	0.1099	0.0062	$0.1 \Omega_{\chi} h^2$	[34]
	Limit (95% CL)		τ (theor.)	ref.
$BR(\overline{B}_s \to \mu^+ \mu^-)$	$< 5.8 imes 10^{-8}$		14%	[35]
m_h	> 114.4 GeV (SM-like Higgs)		$3 \mathrm{GeV}$	[36]
ζ_h^2	$f(m_h)$ (see text)		negligible	[36]
$m_{ ilde{q}}$	$> 375 \mathrm{GeV}$		5%	[25]
$m_{ ilde{g}}$	$> 289 \mathrm{GeV}$		5%	[25]
other sparticle masses	As in table 4 of ref. [6].			

SM parameters

SM (nuisance)	Mean value	Uncertainty	Ref.
parameter	μ	σ (exper.)	
M_t	$172.6{ m GeV}$	$1.4{ m GeV}$	[24]
$m_b(m_b)^{\overline{MS}}$	$4.20{ m GeV}$	$0.07{ m GeV}$	[25]
$\alpha_s(M_Z)^{\overline{MS}}$	0.1176	0.002	[25]
$1/lpha_{ m em}(M_Z)^{\overline{MS}}$	127.955	0.03	[26]

Plus consistency with astrophysical probes

Challenges of profile likelihood evaluation

- MCMC/MultiNest are not designed to find the best-fit point. Bayesian algorithms are designed to map out regions of significant posterior probability mass
- Even for a simple Gaussian toy model, this becomes difficult to do as the number of dimensions of the parameter space increases
- Profiling with vanilla MCMC or MultiNest scans has to be done with caution!

Toy multinormal likelihood

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Posterior pdf from MultiNest scans

- MultiNest is primarily aimed at evaluation of the posterior pdf. It does an excellent job even for multi-modal problems. 8D toy case (Feroz, KC, RT et al, in prep)
- The tolerance parameter (tol) determines the stopping criterium (based on the incremental change of the value of the local evidence). Lower tol gives a finer exploration around the peak, important for profile likelihood evaluation

0

-4

red: analytical

blue: MN, tol=0.5

black: MN, tol=0.0001

0.5

 θ_3

0.5

 θ_3

0.4

0.4

nlive 4000, tol 0.5 nlive 20000, tol 10 4

analytical

0.6

0.6

0.7

0.7

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Profile likelihood from MultiNest scans

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A fairly accurate the profile likelihood can be obtained with MultiNest by tuning the tolerance (lower, tol=0.0001) and the number of live points (higher, n_{live}=20,000) (Feroz, KC, RT et al, in prep), even for highly multi-modal distributions. 8D toy:

red: analytical

blue: MN, tol=0.5

black: MN, tol=0.0001

Profile likelihood from MultiNest

MultiNest scan with 20,000 live points (usually: 4,000) and tolerance 0.0001 (usually: 0.5) results in 5.5 million likelihood evaluations (Akrami et al, GA: 3 million), and best-fit chi-square = 9.26 (Akrami et al, GA: 9.35).

MultiNest finds a better best-fit + smoother contours than GA.

